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CS 551, Spring 2011



- Bayesian Decision Theory is a fundamental statistical approach that quantifies the tradeoffs between various decisions using probabilities and costs that accompany such decisions.
- First, we will assume that all probabilities are known.
- Then, we will study the cases where the probabilistic structure is not completely known.

Fish Sorting Example Revisited

- State of nature is a random variable.
- Define w as the type of fish we observe (state of nature, class) where
 - $w = w_1$ for sea bass,
 - $w = w_2$ for salmon.
 - ▶ $P(w_1)$ is the *a priori probability* that the next fish is a sea bass.
 - $P(w_2)$ is the a priori probability that the next fish is a salmon.

Prior Probabilities

- Prior probabilities reflect our knowledge of how likely each type of fish will appear before we actually see it.
- ▶ How can we choose $P(w_1)$ and $P(w_2)$?
 - ▶ Set $P(w_1) = P(w_2)$ if they are equiprobable (*uniform priors*).
 - May use different values depending on the fishing area, time of the year, etc.
- Assume there are no other types of fish

$$P(w_1) + P(w_2) = 1$$

(exclusivity and exhaustivity).



Making a Decision

► How can we make a decision with only the prior information?

What is the probability of error for this decision?

$$P(error) = \min\{P(w_1), P(w_2)\}\$$



Class-Conditional Probabilities

- ► Let's try to improve the decision using the lightness measurement *x*.
- ▶ Let x be a continuous random variable.
- ▶ Define $p(x|w_j)$ as the *class-conditional probability density* (probability of x given that the state of nature is w_j for j = 1, 2).
- ▶ $p(x|w_1)$ and $p(x|w_2)$ describe the difference in lightness between populations of sea bass and salmon.



Class-Conditional Probabilities

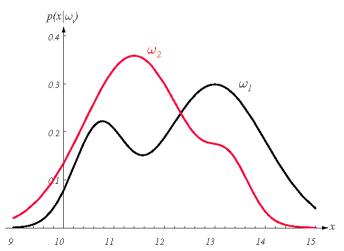


Figure 1: Hypothetical class-conditional probability density functions for two classes.

Posterior Probabilities

- ▶ Suppose we know $P(w_j)$ and $p(x|w_j)$ for j=1,2, and measure the lightness of a fish as the value x.
- ▶ Define $P(w_j|x)$ as the *a posteriori probability* (probability of the state of nature being w_j given the measurement of feature value x).
- ▶ We can use the Bayes formula to convert the prior probability to the posterior probability

$$P(w_j|x) = \frac{p(x|w_j)P(w_j)}{p(x)}$$

where $p(x) = \sum_{j=1}^{2} p(x|w_j)P(w_j)$.



Making a Decision

- ▶ $p(x|w_j)$ is called the *likelihood* and p(x) is called the *evidence*.
- ► How can we make a decision after observing the value of x?

Rewriting the rule gives

$$\text{Decide} \quad \begin{cases} w_1 & \text{if } \frac{p(x|w_1)}{p(x|w_2)} > \frac{P(w_2)}{P(w_1)} \\ w_2 & \text{otherwise} \end{cases}$$

▶ Note that, at every x, $P(w_1|x) + P(w_2|x) = 1$.



Probability of Error

What is the probability of error for this decision?

$$P(error|x) = \begin{cases} P(w_1|x) & \text{if we decide } w_2 \\ P(w_2|x) & \text{if we decide } w_1 \end{cases}$$

What is the average probability of error?

$$P(error) = \int_{-\infty}^{\infty} p(error, x) dx = \int_{-\infty}^{\infty} P(error|x) p(x) dx$$

▶ Bayes decision rule minimizes this error because

$$P(error|x) = \min\{P(w_1|x), P(w_2|x)\}.$$



- ► How can we generalize to
 - more than one feature?
 - ightharpoonup replace the scalar x by the feature vector \mathbf{x}
 - more than two states of nature?
 - just a difference in notation
 - allowing actions other than just decisions?
 - allow the possibility of rejection
 - different risks in the decision?
 - define how costly each action is

- ▶ Let $\{w_1, ..., w_c\}$ be the finite set of c states of nature (*classes*, *categories*).
- ▶ Let $\{\alpha_1, \ldots, \alpha_a\}$ be the finite set of a possible *actions*.
- ▶ Let $\lambda(\alpha_i|w_j)$ be the *loss* incurred for taking action α_i when the state of nature is w_j .
- ► Let x be the *d*-component vector-valued random variable called the *feature vector*.

- $ightharpoonup p(\mathbf{x}|w_i)$ is the class-conditional probability density function.
- ▶ $P(w_i)$ is the prior probability that nature is in state w_i .
- ▶ The posterior probability can be computed as

$$P(w_j|\mathbf{x}) = \frac{p(\mathbf{x}|w_j)P(w_j)}{p(\mathbf{x})}$$

where $p(\mathbf{x}) = \sum_{j=1}^{c} p(\mathbf{x}|w_j) P(w_j)$.



Conditional Risk

- ▶ Suppose we observe x and take action α_i .
- ▶ If the true state of nature is w_j , we incur the loss $\lambda(\alpha_i|w_j)$.
- ▶ The expected loss with taking action α_i is

$$R(\alpha_i|\mathbf{x}) = \sum_{j=1}^{c} \lambda(\alpha_i|w_j) P(w_j|\mathbf{x})$$

which is also called the conditional risk.



Minimum-Risk Classification

- ▶ The general *decision rule* $\alpha(\mathbf{x})$ tells us which action to take for observation \mathbf{x} .
- We want to find the decision rule that minimizes the overall risk

$$R = \int R(\alpha(\mathbf{x})|\mathbf{x}) p(\mathbf{x}) d\mathbf{x}.$$

- ▶ Bayes decision rule minimizes the overall risk by selecting the action α_i for which $R(\alpha_i|\mathbf{x})$ is minimum.
- ► The resulting minimum overall risk is called the *Bayes risk* and is the best performance that can be achieved.

Two-Category Classification

- Define
 - α_1 : deciding w_1 ,
 - α_2 : deciding w_2 ,
 - $\lambda_{ij} = \lambda(\alpha_i|w_j).$
- Conditional risks can be written as

$$R(\alpha_1|\mathbf{x}) = \lambda_{11} P(w_1|\mathbf{x}) + \lambda_{12} P(w_2|\mathbf{x}),$$

$$R(\alpha_2|\mathbf{x}) = \lambda_{21} P(w_1|\mathbf{x}) + \lambda_{22} P(w_2|\mathbf{x}).$$

Two-Category Classification

► The minimum-risk decision rule becomes

lacktriangle This corresponds to deciding w_1 if

$$\frac{p(\mathbf{x}|w_1)}{p(\mathbf{x}|w_2)} > \frac{(\lambda_{12} - \lambda_{22})}{(\lambda_{21} - \lambda_{11})} \frac{P(w_2)}{P(w_1)}$$

 \Rightarrow comparing the *likelihood ratio* to a threshold that is independent of the observation \mathbf{x} .

- \blacktriangleright Actions are decisions on classes (α_i is deciding w_i).
- ▶ If action α_i is taken and the true state of nature is w_j , then the decision is correct if i = j and in error if $i \neq j$.
- We want to find a decision rule that minimizes the probability of error.

Define the zero-one loss function

$$\lambda(\alpha_i|w_j) = \begin{cases} 0 & \text{if } i = j\\ 1 & \text{if } i \neq j \end{cases} \qquad i, j = 1, \dots, c$$

(all errors are equally costly).

Conditional risk becomes

$$R(\alpha_i|\mathbf{x}) = \sum_{j=1}^{c} \lambda(\alpha_i|w_j) P(w_j|\mathbf{x})$$
$$= \sum_{j\neq i} P(w_j|\mathbf{x})$$
$$= 1 - P(w_i|\mathbf{x}).$$

▶ Minimizing the risk requires maximizing $P(w_i|\mathbf{x})$ and results in the *minimum-error decision rule*

Decide
$$w_i$$
 if $P(w_i|\mathbf{x}) > P(w_j|\mathbf{x}) \quad \forall j \neq i$.

► The resulting error is called the *Bayes error* and is the best performance that can be achieved.

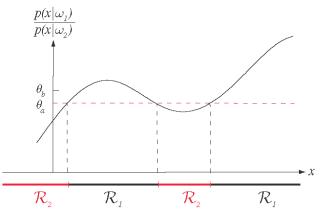


Figure 2: The likelihood ratio $p(\mathbf{x}|w_1)/p(\mathbf{x}|w_2)$. The threshold θ_a is computed using the priors $P(w_1)=2/3$ and $P(w_2)=1/3$, and a zero-one loss function. If we penalize mistakes in classifying w_2 patterns as w_1 more than the converse, we should increase the threshold to θ_b .

Discriminant Functions

▶ A useful way of representing classifiers is through discriminant functions $g_i(\mathbf{x}), i = 1, \dots, c$, where the classifier assigns a feature vector \mathbf{x} to class w_i if

$$g_i(\mathbf{x}) > g_j(\mathbf{x}) \quad \forall j \neq i.$$

► For the classifier that minimizes conditional risk

$$g_i(\mathbf{x}) = -R(\alpha_i|\mathbf{x}).$$

For the classifier that minimizes error

$$g_i(\mathbf{x}) = P(w_i|\mathbf{x}).$$



Discriminant Functions

- ▶ These functions divide the feature space into c decision regions $(\mathcal{R}_1, \ldots, \mathcal{R}_c)$, separated by decision boundaries.
- Note that the results do not change even if we replace every $g_i(\mathbf{x})$ by $f(g_i(\mathbf{x}))$ where $f(\cdot)$ is a monotonically increasing function (e.g., logarithm).
- ► This may lead to significant analytical and computational simplifications.

The Gaussian Density

- Gaussian can be considered as a model where the feature vectors for a given class are continuous-valued, randomly corrupted versions of a single typical or prototype vector.
- ▶ Some properties of the Gaussian:
 - Analytically tractable.
 - Completely specified by the 1st and 2nd moments.
 - Has the maximum entropy of all distributions with a given mean and variance.
 - Many processes are asymptotically Gaussian (Central Limit Theorem).
 - Linear transformations of a Gaussian are also Gaussian.
 - Uncorrelatedness implies independence.



Univariate Gaussian

ightharpoonup For $x \in \mathbb{R}$:

$$p(x) = N(\mu, \sigma^2)$$

$$= \frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{1}{2} \left(\frac{x-\mu}{\sigma}\right)^2\right]$$

where

$$\mu = E[x] = \int_{-\infty}^{\infty} x \, p(x) \, dx,$$

$$\sigma^2 = E[(x - \mu)^2] = \int_{-\infty}^{\infty} (x - \mu)^2 \, p(x) \, dx.$$



Univariate Gaussian

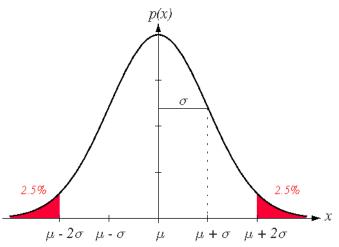


Figure 3: A univariate Gaussian distribution has roughly 95% of its area in the range $|x-\mu| \leq 2\sigma$.

Multivariate Gaussian

ightharpoonup For $\mathbf{x} \in \mathbb{R}^d$:

$$p(\mathbf{x}) = N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$$

$$= \frac{1}{(2\pi)^{d/2} |\boldsymbol{\Sigma}|^{1/2}} \exp \left[-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}) \right]$$

where

$$\boldsymbol{\mu} = E[\mathbf{x}] = \int \mathbf{x} p(\mathbf{x}) d\mathbf{x},$$

$$\boldsymbol{\Sigma} = E[(\mathbf{x} - \boldsymbol{\mu})(\mathbf{x} - \boldsymbol{\mu})^T] = \int (\mathbf{x} - \boldsymbol{\mu})(\mathbf{x} - \boldsymbol{\mu})^T p(\mathbf{x}) d\mathbf{x}.$$



Multivariate Gaussian

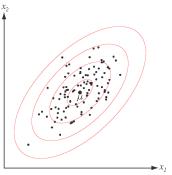


Figure 4: Samples drawn from a two-dimensional Gaussian lie in a cloud centered on the mean μ . The loci of points of constant density are the ellipses for which $(\mathbf{x}-\mu)^T \mathbf{\Sigma}^{-1}(\mathbf{x}-\mu)$ is constant, where the eigenvectors of $\mathbf{\Sigma}$ determine the direction and the corresponding eigenvalues determine the length of the principal axes. The quantity $r^2 = (\mathbf{x}-\mu)^T \mathbf{\Sigma}^{-1}(\mathbf{x}-\mu)$ is called the squared *Mahalanobis distance* from \mathbf{x} to μ .

Linear Transformations

- ► Recall that, given $\mathbf{x} \in \mathbb{R}^d$, $\mathbf{A} \in \mathbb{R}^{d \times k}$, $\mathbf{y} = \mathbf{A}^T \mathbf{x} \in \mathbb{R}^k$, if $x \sim N(\boldsymbol{\mu}, \boldsymbol{\Sigma})$, then $y \sim N(\mathbf{A}^T \boldsymbol{\mu}, \mathbf{A}^T \boldsymbol{\Sigma} \mathbf{A})$.
- ► As a special case, the whitening transform

$$\mathbf{A_w} = \mathbf{\Phi} \mathbf{\Lambda}^{-1/2}$$

where

- Φ is the matrix whose columns are the orthonormal eigenvectors of Σ ,
- lacksquare $lack \Lambda$ is the diagonal matrix of the corresponding eigenvalues,

gives a covariance matrix equal to the identity matrix I.



Discriminant Functions for the Gaussian Density

▶ Discriminant functions for minimum-error-rate classification can be written as

$$g_i(\mathbf{x}) = \ln p(\mathbf{x}|w_i) + \ln P(w_i).$$

 $\qquad \qquad \mathbf{For} \ p(\mathbf{x}|w_i) = N(\boldsymbol{\mu_i}, \boldsymbol{\Sigma_i})$

$$g_i(\mathbf{x}) = -\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu_i})^T \boldsymbol{\Sigma_i}^{-1}(\mathbf{x} - \boldsymbol{\mu_i}) - \frac{d}{2}\ln 2\pi - \frac{1}{2}\ln |\boldsymbol{\Sigma_i}| + \ln P(w_i).$$

Discriminant functions are

$$g_i(\mathbf{x}) = \mathbf{w}_i^T \mathbf{x} + w_{i0}$$
 (linear discriminant)

where

$$\mathbf{w}_{i} = \frac{1}{\sigma^{2}} \boldsymbol{\mu}_{i}$$

$$w_{i0} = -\frac{1}{2\sigma^{2}} \boldsymbol{\mu}_{i}^{T} \boldsymbol{\mu}_{i} + \ln P(w_{i})$$

(w_{i0} is the threshold or bias for the *i*'th category).

▶ Decision boundaries are the hyperplanes $g_i(\mathbf{x}) = g_j(\mathbf{x})$, and can be written as

$$\mathbf{w}^T(\mathbf{x} - \mathbf{x_0}) = 0$$

where

$$\begin{split} \mathbf{w} &= \boldsymbol{\mu_i} - \boldsymbol{\mu_j} \\ \mathbf{x_0} &= \frac{1}{2}(\boldsymbol{\mu_i} + \boldsymbol{\mu_j}) - \frac{\sigma^2}{\|\boldsymbol{\mu_i} - \boldsymbol{\mu_j}\|^2} \ln \frac{P(w_i)}{P(w_j)} (\boldsymbol{\mu_i} - \boldsymbol{\mu_j}). \end{split}$$

▶ Hyperplane separating \mathcal{R}_i and \mathcal{R}_j passes through the point $\mathbf{x_0}$ and is orthogonal to the vector \mathbf{w} .

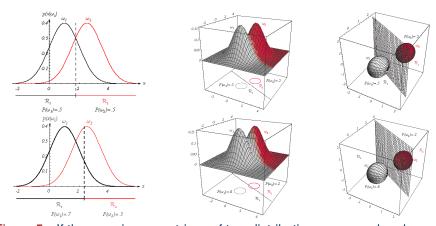


Figure 5: If the covariance matrices of two distributions are equal and proportional to the identity matrix, then the distributions are spherical in d dimensions, and the boundary is a generalized hyperplane of d-1 dimensions, perpendicular to the line separating the means. The decision boundary shifts as the priors are changed.

▶ Special case when $P(w_i)$ are the same for i = 1, ..., c is the minimum-distance classifier that uses the decision rule

assign
$$\mathbf{x}$$
 to w_{i^*} where $i^* = \arg\min_{i=1,\dots,c} \|\mathbf{x} - \boldsymbol{\mu_i}\|$.

Case 2: $\Sigma_i = \Sigma$

Discriminant functions are

$$g_i(\mathbf{x}) = \mathbf{w}_i^T \mathbf{x} + w_{i0}$$
 (linear discriminant)

where

$$\mathbf{w}_{i} = \mathbf{\Sigma}^{-1} \boldsymbol{\mu}_{i}$$

$$w_{i0} = -\frac{1}{2} \boldsymbol{\mu}_{i}^{T} \mathbf{\Sigma}^{-1} \boldsymbol{\mu}_{i} + \ln P(w_{i}).$$

Case 2: $\Sigma_i = \Sigma$

Decision boundaries can be written as

$$\mathbf{w}^T(\mathbf{x} - \mathbf{x_0}) = 0$$

where

$$\begin{split} \mathbf{w} &= \mathbf{\Sigma}^{-1}(\boldsymbol{\mu_i} - \boldsymbol{\mu_j}) \\ \mathbf{x_0} &= \frac{1}{2}(\boldsymbol{\mu_i} + \boldsymbol{\mu_j}) - \frac{\ln(P(w_i)/P(w_j))}{(\boldsymbol{\mu_i} - \boldsymbol{\mu_j})^T \mathbf{\Sigma}^{-1}(\boldsymbol{\mu_i} - \boldsymbol{\mu_j})} (\boldsymbol{\mu_i} - \boldsymbol{\mu_j}). \end{split}$$

Hyperplane passes through x₀ but is not necessarily orthogonal to the line between the means.



Case 2: $\Sigma_i = \Sigma$

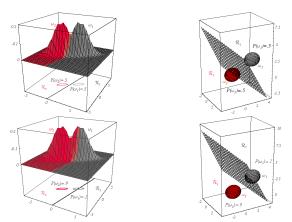


Figure 6: Probability densities with equal but asymmetric Gaussian distributions. The decision hyperplanes are not necessarily perpendicular to the line connecting the means.

Case 3: Σ_i = arbitrary

Discriminant functions are

$$g_i(\mathbf{x}) = \mathbf{x}^T \mathbf{W}_i \mathbf{x} + \mathbf{w}_i^T \mathbf{x} + w_{i0}$$
 (quadratic discriminant)

where

$$\begin{aligned} \mathbf{W}_{i} &= -\frac{1}{2} \boldsymbol{\Sigma}_{i}^{-1} \\ \mathbf{w}_{i} &= \boldsymbol{\Sigma}_{i}^{-1} \boldsymbol{\mu}_{i} \\ w_{i0} &= -\frac{1}{2} \boldsymbol{\mu}_{i}^{T} \boldsymbol{\Sigma}_{i}^{-1} \boldsymbol{\mu}_{i} - \frac{1}{2} \ln |\boldsymbol{\Sigma}_{i}| + \ln P(w_{i}). \end{aligned}$$

Decision boundaries are hyperquadrics.



Case 3: Σ_i = arbitrary

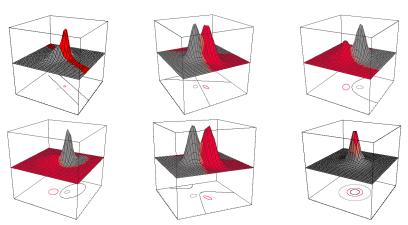


Figure 7: Arbitrary Gaussian distributions lead to Bayes decision boundaries that are general hyperquadrics.

Case 3: Σ_i = arbitrary

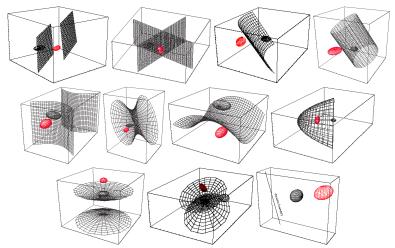


Figure 8: Arbitrary Gaussian distributions lead to Bayes decision boundaries that are general hyperquadrics.

Error Probabilities and Integrals

► For the two-category case

$$P(error) = P(\mathbf{x} \in \mathcal{R}_2, w_1) + P(\mathbf{x} \in \mathcal{R}_1, w_2)$$

$$= P(\mathbf{x} \in \mathcal{R}_2 | w_1) P(w_1) + P(\mathbf{x} \in \mathcal{R}_1 | w_2) P(w_2)$$

$$= \int_{\mathcal{R}_2} p(\mathbf{x} | w_1) P(w_1) d\mathbf{x} + \int_{\mathcal{R}_1} p(\mathbf{x} | w_2) P(w_2) d\mathbf{x}.$$

Error Probabilities and Integrals

► For the multicategory case

$$P(error) = 1 - P(correct)$$

$$= 1 - \sum_{i=1}^{c} P(\mathbf{x} \in \mathcal{R}_i, w_i)$$

$$= 1 - \sum_{i=1}^{c} P(\mathbf{x} \in \mathcal{R}_i | w_i) P(w_i)$$

$$= 1 - \sum_{i=1}^{c} \int_{\mathcal{R}_i} p(\mathbf{x} | w_i) P(w_i) d\mathbf{x}.$$

Error Probabilities and Integrals

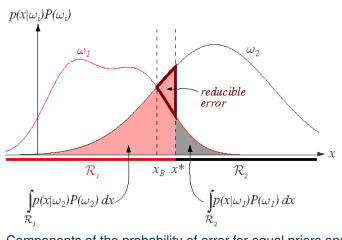


Figure 9: Components of the probability of error for equal priors and the non-optimal decision point x^* . The optimal point x_B minimizes the total shaded area and gives the Bayes error rate.

Receiver Operating Characteristics

- Consider the two-category case and define
 - w_1 : target is present,
 - w_2 : target is not present.

Table 1: Confusion matrix.

		Assigned	
		w_1	w_2
True	w_1	correct detection	mis-detection
	w_2	false alarm	correct rejection

- ▶ Mis-detection is also called false negative or Type I error.
- ► False alarm is also called false positive or Type II error.

Receiver Operating Characteristics

► If we use a parameter (e.g., a threshold) in our decision, the plot of these rates for different values of the parameter is called the receiver operating characteristic (ROC) curve.

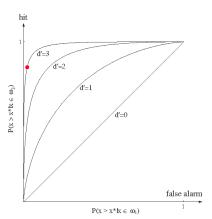


Figure 10: Example receiver operating characteristic (ROC) curves for different settings of the system.

Summary

- ▶ To minimize the overall risk, choose the action that minimizes the conditional risk $R(\alpha|\mathbf{x})$.
- ► To minimize the probability of error, choose the class that maximizes the posterior probability $P(w_i|\mathbf{x})$.
- ▶ If there are different penalties for misclassifying patterns from different classes, the posteriors must be weighted according to such penalties before taking action.
- ▶ Do not forget that these decisions are the optimal ones under the assumption that the "true" values of the probabilities are known.

